Assessment of post-trauma complications in eight million trauma cases over a decade in the USA

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ABSTRACT

Background Trauma is associated with a significant risk of post-trauma complications (PTCs). These include thromboembolic events, strokes, infections, and failure of organ systems (eg, kidney failure). Although care of the trauma patient has evolved during the last decade, whether this has resulted in a reduction in specific PTCs is unknown. We hypothesize that the incidence of PTCs has been decreasing during a 10-year period from 2007 to 2017.

Methods This is a descriptive study of trauma patients originating from level 1, 2, 3, and 4 trauma centers in the USA, obtained via the Trauma Quality Improvement Program (TQIP) database from 2007 to 2017. PTCs documented throughout the time frame were extracted along with demographic variables. Multiple regression modeling was used to associate admission year with PTCs, while controlling for age, gender, Glasgow Coma Scale score, and Injury Severity Score.

Results Data from 8 720 026 trauma patients were extracted from the TQIP database. A total of 366 768 patients experienced one or more PTCs. There was a general decrease in the incidence of PTCs during the study period, with the overall incidence dropping from 7.0% in 2007 to 2.8% in 2017. Multiple regression identified a slight decrease in incidence in all PTCs, although deep surgical site infection (SSI), deep venous thrombosis (DVT), and stroke incidences increased when controlled for confounders.

Discussion Overall the incidence of PTCs dropped during the 10-year study period, although deep SSI, DVT, stroke, and cardiac arrest increased during the study period. Better risk prediction tools, enabling a precision medicine approach, are warranted to identify at-risk patients.

Level of evidence III.

INTRODUCTION

Although major trauma survival rates have improved during the last decades worldwide,¹ survivors still face a significant risk of morbidity (post-trauma complications, PTCs). These include thromboembolic events, infections, stroke, organ failure, and sepsis. Previous reports have indicated that almost half of all trauma patients require intensive care unit (ICU) stays, while upwards of 23% of these cases experience a PTC.² This is thus a significant addition to trauma-related morbidity, further underlined by reports indicating that each hospital complication increases the Odds ratio (OR) for hospital mortality by 2.3.² For trauma patients admitted to the ICU, a previous study has reported

an increase in mortality rate from 10.7% to 16.9% if patients experienced a PTC.²

Previous studies have sought to address both prehospital as well as in-hospital risk factors,^{3–15} including fielding models for early in-hospital prediction of PTC risk.¹⁶ As such, numerous improvements in the care of the injured during the last decade, including adherence to thromboprophylaxis and respiratory protocols, as well as optimal fluid resuscitation strategies, could theoretically translate into a gradual reduction in the incidence of PTCs over time. Whether such reductions can be observed during the last decade is currently unknown and constitutes the primary focus of this study. Despite these improvements in treatment protocols and standards, a number of patients may still be at risk of developing PTCs, owing to factors such as genetic composition, comorbidities, lifestyle choices, and so on. The size of this cohort and thus the scope of the problem of trauma patients unresponsive to current prophylaxis protocols is unknown. Investigating how many patients could potentially benefit from novel, precision medicine-based approaches in PTC prevention presents the secondary aim of this study.

We hypothesized that the incidence of PTCs has reduced during the 10-year study period from 2007 to 2017, but that a number of patients would still suffer these PTCs and thus be potential candidates for future precision medicine-based approaches.

PATIENTS AND METHODS

Access to the Trauma Quality Improvement Program (TQIP) database was granted by the American College of Surgeons TQIP, and data were accessed and handled in line with the TQIP data user agreement. We extracted the PTCs that were available for all years during the study time frame. Online supplemental table 1 lists the chosen 12 PTCs as well as their definitions, as defined by TQIP. We furthermore extracted demographic and injury characteristics variables, including data on age, gender, Glasgow Coma Scale (GCS), Injury Severity Score (ISS), Abbreviated Injury Scale (AIS), disposition on discharge from the emergency department, hospital length of stay (LOS), and time on ventilator. For one variable (pneumonia), values for 2016 and 2017 were excluded due to changes in the underlying data definition compared with previous years. Incidences of the selected PTCs were calculated during the study years.

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Regression models

To assess whether a significant change in rates of PTCs had occurred, we constructed logistic regression models using the occurrence of the given PTC as the dependent variable, with admission year serving as the predictor. We present univariate model predictions as well as multiple regression modeling, correcting for patients' age, gender, GCS score, and ISS. Models were furthermore controlled for the presence of traumatic brain injury (TBI, defined as a head AIS score of >2 as previously suggested as a definition for moderate to severe TBI¹⁷), early mortality (defined as death in the emergency department), LOS, and time on ventilator.

The rationale for including early mortality and LOS in the model hinged on the fact that early mortality would preclude patients from developing a range of complications, whereas early discharge would mean that the patient was lost to follow-up. Furthermore, the model was controlled for duration of ventilator treatment (where applicable) as this factor is well known to precipitate pulmonary complications.

The selected confounding variables were chosen after appraisal of variables with a perceived significant impact on PTC incidence. Shock-related variables (eg, lactate levels and/or base excess) were assessed but were not consistently available in the data set.

Missing data

From 2007 to 2016, TQIP registered PTCs in a separate complications table, linked to the main table with an identification key. For each trauma entry in the main table, one or more entries exist in the complications table, indicating either missing data, no complication or a number of entries corresponding to the number of PTCs recorded. If data were recorded as missing, it was thus not possible to dissect which specific PTC was missing, solely that complication data were missing in total. This approach was changed for the 2017 TQIP data set, where each complication was coded separately. To test the effect of the missing data, a sensitivity analysis was performed. To this end, we created an imputed data set using predictive mean matching as implemented in the R MICE package.¹⁸ The regression models were then applied to the imputed data set for comparative purposes.

Data presentation

Data are presented as median (IQR) for continuous variables and percentages for dichotomous variables, where appropriate. Results of regression modeling are presented as OR with 95% CI. A p value <0.05 was considered statistically significant. Statistical analyses were performed using R.¹⁹

RESULTS

A total of 8 720 026 trauma patients were identified during the 10-year study period, with inclusion rates rising from 506 257 in 2007 to 997 970 in 2017. Demographic characteristics and outcomes are summarized in [table 1](#). In summary, 5 402 999 (61.9%) were male, 3 293 987 (37.8%) were female, and 23 040 (0.3%) did not have a registered gender. The median age was 43 (23–64) years. The overall mortality rate was 3.6% (331 956 patients).

A total of 366 768 (4.2%) patients experienced a PTC. Of these, 78% experienced one complication, 16% two complications, and 4.3% three complications, eclipsing eleven PTCs for the most critical cases.

[Table 2](#) lists the incidence of complications throughout the study period, and the development over time is graphically

Table 1 Demographic and clinical data

Age, years	43 (23–64)
Male, n (%)	5 402 999 (62.1)
Survivors, n (%)	8 388 070 (96.2)
ISS	8 (4–10)
GCS	14 (3–15)
PTC, n (%)	366 768 (4.2)
Number of PTCs, n (%)	
0	8 353 258 (95.8)
1	286 252 (3.2)
2	58 698 (0.67)
3	15 861 (0.18)
4	4379 (0.050)
5	1161 (0.013)
6	327 (0.0038)
7	78 (0.00089)
8	8 (0.000092)
9	3 (0.000034)
11	1 (0.000011)

Data are presented as median (IQR) or percentage where appropriate.

GCS, Glasgow Coma Scale; ISS, Injury Severity Score; PTC, post-trauma complication.

depicted in [figure 1](#). Overall, the incidence of PTC decreased from 7.0% in 2007 to 2.8% in 2017.

When assessing unadjusted incidence rates, the most common complication throughout all years was pneumonia, with an incidence of between 2.1% in 2007 and 1.5% in 2015 of PTCs ([table 2](#)). Acute respiratory distress syndrome (ARDS) declined from an incidence of 1.2% to 0.3%, deep venous thrombosis (DVT)/thrombophlebitis from 0.7% to 0.5%, acute renal failure from 0.7% to 0.4%, cardiac arrest unchanged from 0.6% to 0.6%, superficial surgical site infection (SSI) from 0.2% to 0.1%, and sepsis from 0.7% to 0.3%. In contrast, stroke saw a general increase in incidence through the entire period, from 0.1% in 2007 to 0.2% in 2017.

Regression models

The results of the regression models are shown in [table 3](#). Multivariate modeling confirmed a significant decrease over time for PTCs, including acute renal failure (OR 0.97, CI 0.96 to 0.97), ARDS (OR 0.88, CI 0.88 to 0.89), myocardial infarction (OR 0.97, CI 0.97 to 0.98), cardiac arrest (OR 0.89, CI 0.88 to 0.89), organ space infection (OR 0.97, CI 0.97 to 0.98), pneumonia (OR 0.97, CI 0.97 to 0.98), superficial SSI (OR 0.97, CI 0.97 to 0.97), and systemic sepsis (OR 0.67, CI 0.67 to 0.68). In contrast, a significant increase over time in PTCs including deep SSI (OR 1.08, CI 1.07 to 1.10), DVT/thrombophlebitis (OR 1.09, CI 1.08 to 1.09), and stroke (OR 1.05, CI 1.05 to 1.06) was identified.

Missing data

Online supplemental table 2 provides an overview of the missing data. Overall, 5.8% of records had one or more variables missing. Online supplemental table 3 lists the results of the regression models when applied to the imputed data set. Although ORs differed between the raw and imputed data sets ([table 3](#) and online supplemental table 3), the directionality of these did not differ. As such, sensitivity analysis did not indicate a substantial impact of the missing data on the regression model results.

Table 2 Complication categories by year

	2007	2008	2009	2010	2011	2012	2013	2014	2015	2016	2017
Patients reported (n)	506 257	633 952	684 460	721 536	783 476	830 774	827 289	858 362	912 816	963 134	997 970
Acute renal failure	3523 (0.7)	4568 (0.7)	5181 (0.7)	5347 (0.7)	5479 (0.7)	5555 (0.7)	5687 (0.7)	4249 (0.5)	4738 (0.5)	4814 (0.5)	4411 (0.4)
ARDS	6551 (1.2)	7198 (1.1)	7530 (1.1)	7766 (1.0)	9743 (1.2)	7863 (0.9)	6357 (0.7)	4146 (0.5)	3404 (0.4)	3210 (0.3)	2634 (0.3)
Cardiac arrest with CPR	2824 (0.6)	4274 (0.7)	4453 (0.6)	4561 (0.6)	5143 (0.6)	5694 (0.7)	5366 (0.6)	5140 (0.6)	6295 (0.7)	6041 (0.6)	6086 (0.6)
Deep SSI	111 (0.0)	261 (0.0)	349 (0.1)	530 (0.1)	633 (0.1)	805 (0.1)	796 (0.1)	901 (0.1)	974 (0.1)	1090 (0.1)	882 (0.1)
DVT/thrombophlebitis	3688 (0.7)	4884 (0.7)	5068 (0.7)	4209 (0.6)	5792 (0.7)	5913 (0.7)	5707 (0.7)	5526 (0.6)	5386 (0.6)	5194 (0.5)	4766 (0.5)
Myocardial infarction	1093 (0.2)	1471 (0.2)	1415 (0.2)	1465 (0.2)	1879 (0.2)	2008 (0.2)	2027 (0.2)	1663 (0.2)	1627 (0.2)	1514 (0.2)	1449 (0.1)
Organ/space SSI	938 (0.2)	1163 (0.2)	1382 (0.2)	1394 (0.2)	1352 (0.2)	2689 (0.2)	2485 (0.1)	1139 (0.1)	1131 (0.1)	798 (0.1)	788 (0.1)
Pneumonia	10 827 (2.1)	14 062 (2.1)	15 133 (2.1)	15 211 (2.0)	16 379 (2.0)	16 909 (2.0)	15 547 (1.8)	14 717 (1.6)	14 214 (1.5)	N/A	N/A
PE	1492 (0.3)	2160 (0.3)	2141 (0.3)	2157 (0.3)	2191 (0.3)	2457 (0.3)	2276 (0.3)	2341 (0.3)	2422 (0.3)	2440 (0.2)	2409 (0.2)
Stroke/CVA	483 (0.1)	737 (0.1)	953 (0.1)	1019 (0.1)	1049 (0.1)	1898 (0.2)	2026 (0.2)	1778 (0.2)	1817 (0.2)	1948 (0.2)	2016 (0.2)
Superficial SSI	647 (0.2)	849 (0.1)	1109 (0.2)	1124 (0.1)	1153 (0.1)	1324 (0.2)	N/A	1138 (0.1)	1314 (0.1)	1065 (0.1)	776 (0.1)
Systemic sepsis	3489 (0.7)	4004 (0.6)	4003 (0.5)	3950 (0.5)	N/A	2305 (0.3)	2603 (0.3)	2797 (0.3)	3109 (0.3)	2564 (0.3)	2507 (0.3)

Numbers indicate the total number of patients for the year in question (incidence %).

Pneumonia was not included for 2016 and 2017 due to changes in data definition.

Data on superficial SSI and sepsis were not available for years 2013 and 2011, respectively.

ARDS, acute respiratory distress syndrome; CPR, cardiopulmonary resuscitation; CVA, cerebrovascular accident; DVT, deep venous thrombosis; N/A, not available; PE, pulmonary embolism; SSI, surgical site infection.

Figure 1 Graphical overview of the development of post-trauma complications from 2007 to 2017. CPR, cardiopulmonary resuscitation; CVA, cerebrovascular accident.

DISCUSSION

In this study, we assessed the development of PTCs in the USA during a 10-year period from 2007 to 2017. We hypothesized that temporal PTC incidences would have dropped during the investigative time frame, but that a number of patients remain at risk. Of the 12 selected PTC categories, we did indeed identify significant reductions in the PTC incidence of acute renal failure, ARDS, myocardial infarction, organ space and superficial SSI, pneumonia, pulmonary embolism, cardiac arrest, and sepsis. In contrast, the incidence of deep SSI, DVT, and stroke increased over time when multivariate modeling was considered. Although the graphical representation in [figure 1](#) suggests an overall drop in PTC incidence, multivariate modeling thus indicates a slight increase in certain PTCs, which runs counter to the study hypothesis.

In the last study year, where treatment protocols would presumably be optimal compared with the previous years, 27 943 (2.8%) patients experienced one or more PTCs. To put this number into perspective, the global burden of injury in the 2017 study estimated that 520 million patients worldwide suffered traumatic injuries in that year.²⁰ If numbers can be extrapolated, this would mean that an excess of 14 million patients could have suffered a PTC globally in 2017 alone. Although these numbers should be interpreted in the light of the many differences in trauma systems worldwide, they still suggest that a sizeable number of patients remain unresponsive to current prophylaxis protocols and could potentially benefit from precision medicine-based approaches.

Several factors should be considered when interpreting the data, and caution should be exercised when interpreting the presented results. The annual volume of trauma patients in the TQIP data set increased from 2007 (506 257) to 2017 (997 970), and the underlying demographics and injury characteristics also varied over the years. The fluctuations in complication rates could

Table 3 Results of regression model using the complication in question as the dependent variable and the admission year as the predictor

	Corrected model			Univariate model		
	OR	95% CI	P value	OR	95% CI	P value
Acute renal failure	0.97	0.96 to 0.97	<0.01	0.93	0.93 to 0.94	<0.01
ARDS	0.89	0.88 to 0.89	<0.01	0.85	0.84 to 0.85	<0.01
Cardiac arrest with CPR	0.89	0.88 to 0.89	<0.01	0.99	0.99 to 0.99	<0.01
Deep SSI	1.08	1.07 to 1.10	<0.01	1.09	1.08 to 1.10	<0.01
DVT/thrombophlebitis	1.09	1.08 to 1.09	<0.01	0.96	0.96 to 0.96	<0.01
Myocardial infarction	0.97	0.97 to 0.98	<0.01	0.95	0.94 to 0.95	<0.01
Organ/space SSI	0.97	0.97 to 0.98	<0.01	0.87	0.86 to 0.87	<0.01
Pneumonia	0.97	0.97 to 0.98	<0.01	0.86	0.86 to 0.87	<0.01
Pulmonary embolism	0.97	0.96 to 0.97	<0.01	0.96	0.95 to 0.96	<0.01
Stroke/CVA	1.01	1.01 to 1.02	<0.01	1.04	1.04 to 1.05	<0.01
Superficial SSI	0.97	0.97 to 0.97	<0.01	0.95	0.94 to 0.95	<0.01
Systemic sepsis	0.67	0.67 to 0.68	<0.01	0.71	0.70 to 0.72	<0.01

Data are presented as OR with 95% CI and p values.

The multivariate model was controlled for confounders, including age, gender, injury severity, and Glasgow Coma Scale scores.

ARDS, acute respiratory distress syndrome; CPR, cardiopulmonary resuscitation; CVA, cerebrovascular accident; DVT, deep venous thrombosis; SSI, surgical site infection.

thus in part be due to data from additional trauma centers with demographic and injury characteristics variations being added to the data set. As treatment protocol adjustment after feedback from the TQIP would take time to implement, the rapid increase in the number of participating centers may thus create a setting where centers would enter TQIP with suboptimal PTC incidence rates, which would be gradually corrected once TQIP feedback was obtained and time for protocol optimization was allowed for. Furthermore, as is the case for many retrospective databases, there is likely a significant issue of under-reporting in the data set. As such, a recent study from Japan, investigating 184 214 patients, reported a PTC rate of 12.8%,²¹ as opposed to 2.8% in this study. Although obvious differences in number and definitions of reported PTC exist between data sets, the presented results should be interpreted in light of the underlying data set. Although TQIP likely also suffers from under-reporting, it is less clear whether such an under-reporting should exhibit a temporal trend. As such, it is likely that the findings of a relatively stable PTC trend for most complications reflect reality, although at a higher incidence than reported here.

Also, the results should be interpreted in the light of the inherent variance in reporting standards between sites this and most other retrospective quality registers suffer from. Indeed, studies have indicated that a degree of interobserver variability exists in TQIP, which could affect the presented results.²²

The observed reduction in the incidence of pulmonary complications, including ARDS and pneumonia, can likely be associated with the development and adherence to resuscitation and ventilator-associated pneumonia (VAP) protocols during the last decade, including outcomes of research collaboratives such as the ARDS Network.²³ High incidences of pneumonia are well documented as a major cause of PTC,² and VAP continues to be a preventable burden for critically ill patients. Studies on prevention strategies have shown variable success, focusing on treatments including non-invasive positive pressure ventilation, optimal bed position, better oral care, and removal of subglottic secretions.^{24 25} Hospitals adhering to ventilator optimization strategies have reported good drops in incidence statistics.²⁶

ARDS incidences are likely also affected by developments in resuscitation strategies, as well as accelerated patient mobilization efforts.²⁷ As such, the gradual shift from large-volume crystalloid resuscitation toward a permissive hypotensive and

balanced blood transfusion regimen has likely played a role in reducing ARDS incidences^{28–30} nationwide.

Acute renal failure, associated with increased morbidity and mortality as well as hospital LOS,³¹ also exhibited reduced incidence during the study period. This is likely associated with the development of and adherence to risk assessment protocols targeting renal failure, mainly through optimizing renal perfusion.³² Of note, the second most common reason for renal failure is sepsis, which also exhibited reduced incidence during the study period.

Collectively, it is likely that these improvements are due to increased adherence to updated resuscitation and treatment protocols, including sepsis, as well as resuscitation and treatment guidelines such as those championed by trauma societies.^{33 34}

For thrombosis ORs, we observed a decrease in pulmonary embolism, but a slight increase in DVT and stroke ORs over time when multivariate modeling was considered. Although thrombosis prophylaxis protocols have received much attention,³⁵ with an apparent drop in unadjusted DVT rates over time, multivariate modeling suggested that this could be due to changes in the underlying demographics and trauma characteristics of included patients. As such, current prophylaxis protocols have been unsuccessful in further reducing DVT and stroke incidences over time when patient covariates are considered, thus highlighting a focus area for further research and development.

Although myocardial infarction and cardiac arrest incidence rates decreased in this study, the observed trend for DVT and stroke thus mirrors the rise in cardiovascular disease-related mortality in the general US population during the last decade³⁶ and could potentially be related to changes in lifestyle factors, including diet, smoking, and a general increase in sedentary lifestyle.³⁶ For DVT, the observed increase in adjusted incidence is in line with a general embolism-associated mortality increase in USA since 2008.^{37 38} Interestingly, this did not translate into an increased incidence of pulmonary embolism in this study, which could potentially be associated with focus on vena cava filters in high-risk patients,³⁹ although this cannot be concluded from these data.

For infectious complications, we observed a reduction in ORs of organ space and superficial SSI, although an increase in deep SSI OR was observed concurrently. Whether this represents real fluctuations associated with changes in treatment strategies (eg,

increase in non-operative management strategies) or simply a shift in SSI classification practices cannot be readily deduced from these data. Incidences of SSI have in other studies decreased and were largely associated with small bowel and vascular bypass surgery.^{40,41} The reduction in SSI incidence observed here is thus in line with previous reports from non-trauma surgical patients. These results should, however, be interpreted with caution. The structure of the TQIP database did not allow for a consistent registration of the nature, indication, and type of surgical procedures. As such, fluctuations in the number of major surgical procedures performed during the investigative time frame could have impacted on the results.

Overall, although selected PTC incidences have shown a temporal decrease (pneumonia and ARDS), other PTCs have failed to show a clear development. Although these exhibited slight increases or reductions, it is questionable whether the magnitude of these fluctuations is of clinical relevance (figure 1). Furthermore, although most of these alterations are statistically significant, this should be analyzed in light of the large number of patients present for analysis.

The study has several limitations. First, this is a retrospective study dependent on the quality and correctness of data sources from the TQIP database. As such, PTCs may have been missed by the curators. Second, although we have sought to control for relevant confounders, the results may still be affected by factors not included in the regression model. Such factors could potentially include the number of major surgical procedures performed, as this could have impacted on PTC incidences, specifically Venous Thromboembolisms (VTE) and SSI rates. Third, data from the TQIP database are limited to trauma centers participating in the program, which may not completely mirror other centers throughout the USA or elsewhere. Fourth, although the sensitivity analysis did not indicate a major effect of the missing data, interpretation of the presented results should be seen in light of the fact that TQIP data quality and data completeness generally increased toward the final years of the study period. The increase in the number of participating hospitals could also have affected data quality as well as reported outcomes, either due to variations in data definitions or differential outcomes between participating centers. It would thus have been interesting to identify hospitals present in TQIP throughout the study time frame to assess PTC variations in these. TQIP, does, however, not allow for an identification of the individual center, which precludes us from making this analysis. Also, the use of advanced directives could have impacted on the level of treatment offered to patients. The TQIP data set does, however, only contain information on such decisions from 2013 and onwards, which was considered incompatible with the analysis approach.

The TQIP data structure was changed during the study period, with the number and types of PTCs recorded differing from 2007 to 2017. Ideally, a uniform data set would have been optimal, and fluctuations in data definitions and recorded variables could thus impact on the results. The TQIP data set adheres to the definition standard set forth by the National Trauma Data Bank, as defined in the National Trauma Data Set (NTDS) standards. There are annual updates of the PTC definitions, and variations could thus affect the presented results. A review of the NTDS changelogs did, however, reflect minor changes with perceived limited impact on the presented findings. For comparative purposes, we provide an overview of the 2007 vs 2017 PTC definitions in online supplemental table 1.

Finally, certain PTCs such as venous embolisms are critically dependent on imaging studies for their detection. The TQIP data set does not allow for an assessment of the use of imaging

modalities. As such, whether changes in the frequency of imaging studies could have impacted on the presented findings cannot be deduced from these data but may impact on the results. Collectively, interpretation of the presented results should thus be done with these limitations in mind.

Even with these limitations, we conclude that incidences of PTC remain largely stationary over time, with a slight decrease or increase for selected PTCs. A number of patients remain unresponsive to current treatment prophylaxis and could be candidates for future precision medicine-based approaches.

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Patient consent for publication Not required.

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Data availability statement Data may be obtained from a third party and are not publicly available. Data are available upon submission of a research proposal to the TQIP administrators.

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